

# Tensor Moments of Gaussian Mixture Models: Theory and Applications

Tammy Kolda MathSci.ai

Joe Kileel and João M. Pereira University of Texas, Austin

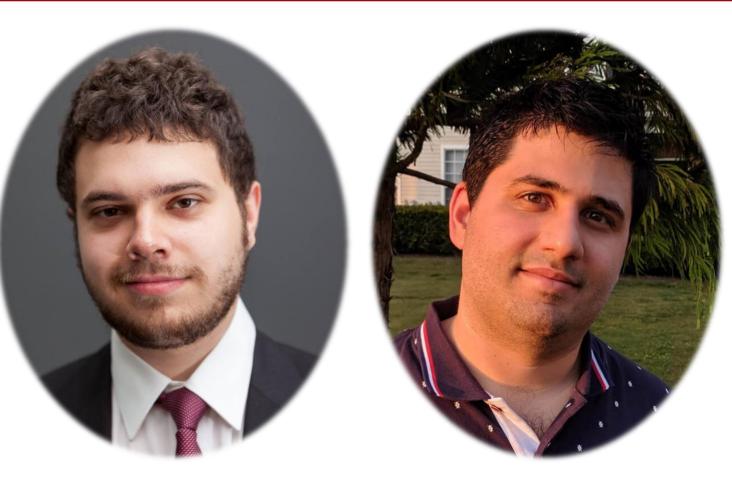
Chris Brigm

Kolda @ Texas A&M TRIPODS Distinguished Lecture

#### **Amazing Coauthors**



Joe Kileel Asst Professor Math Dept U Texas, Austin



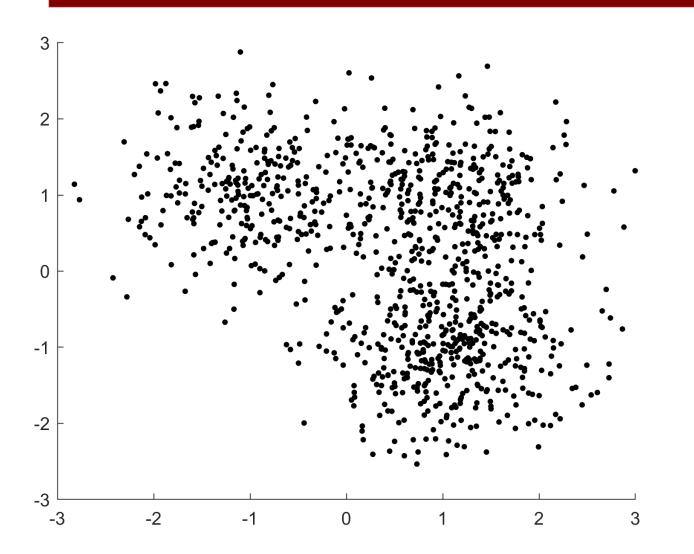
João M. Pereira Postdoc Math Dept U Texas, Austin

(starting summer 2022) Asst Professor Instituto Nacional de Matemática Pura e Aplicada (IMPA) Brazil



# Motivation: Making Sense of Data via Models

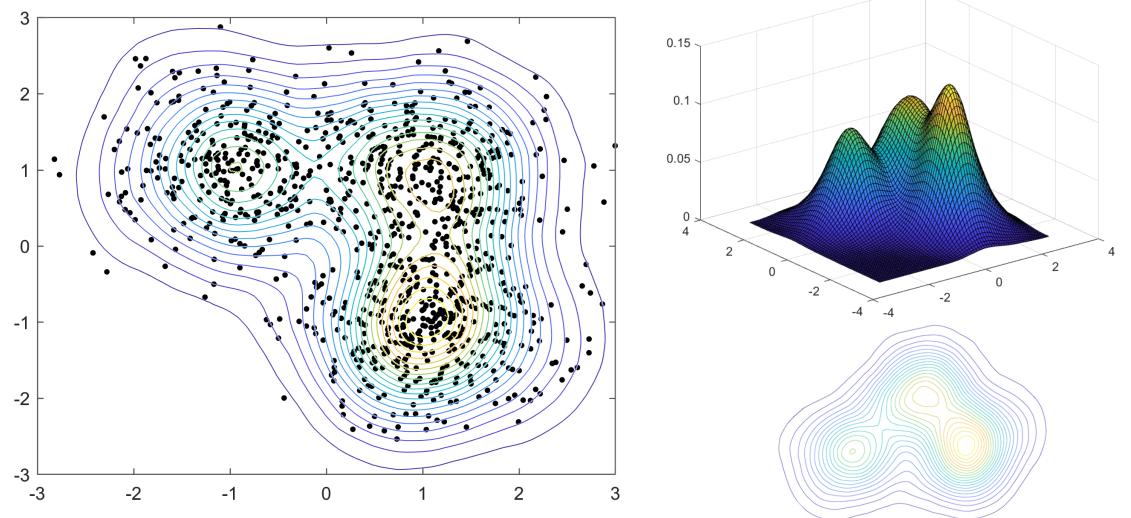
#### Motivation: Making Sense of Data



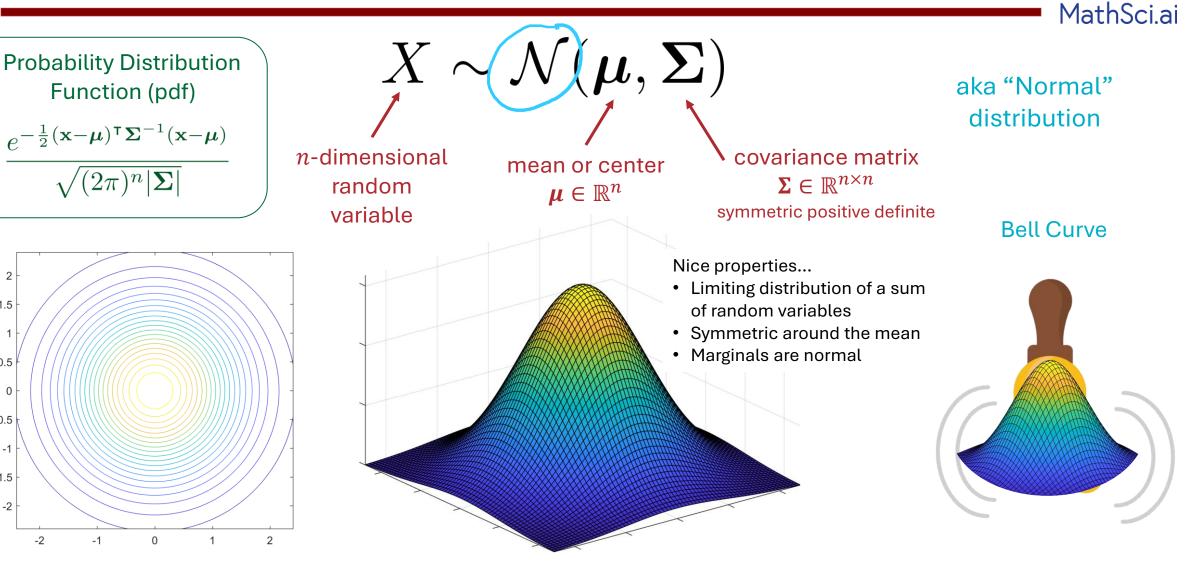
Illustrative Example: 1000 datapoints  $(x_1, x_2)$  MathSci.ai

#### Technique: Create Model of Probability Density from Datapoints





## Gaussian Distribution ("Bell curve")



-2

2

1.5

0.5

0

-0.5

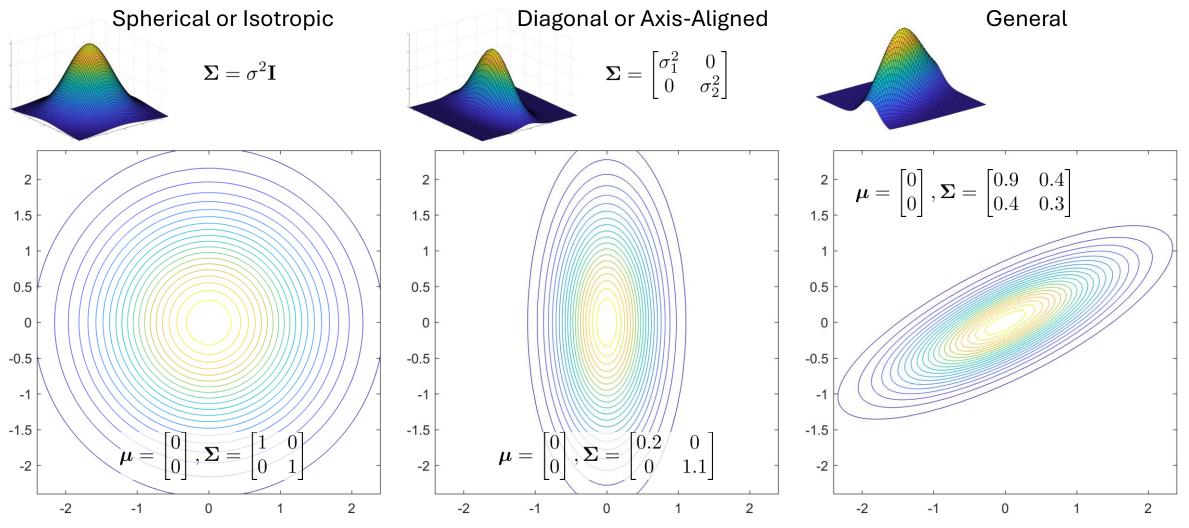
-1

-1.5

-2

## **Covariance for Gaussian Distribution**

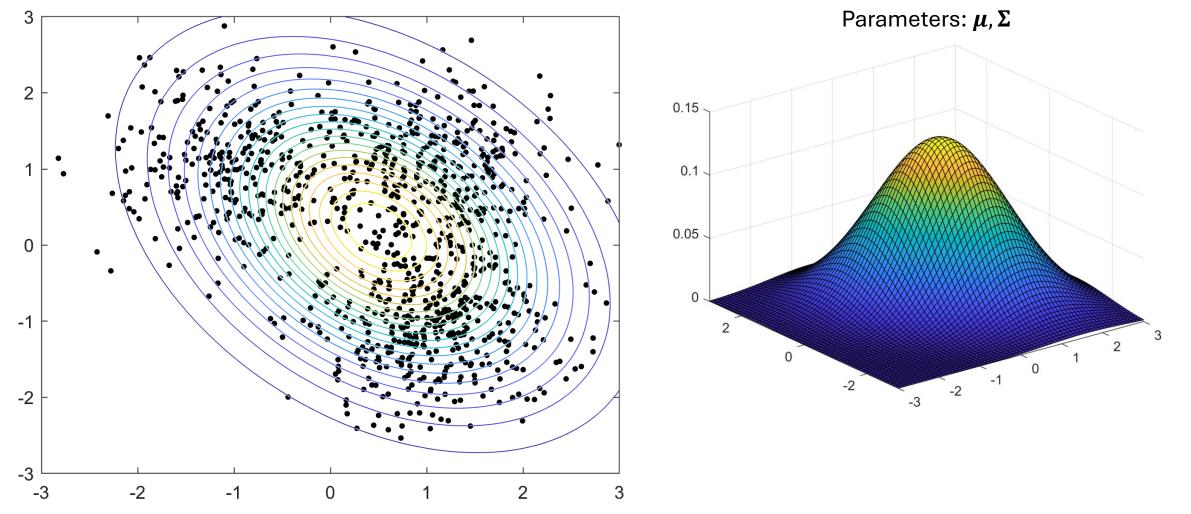




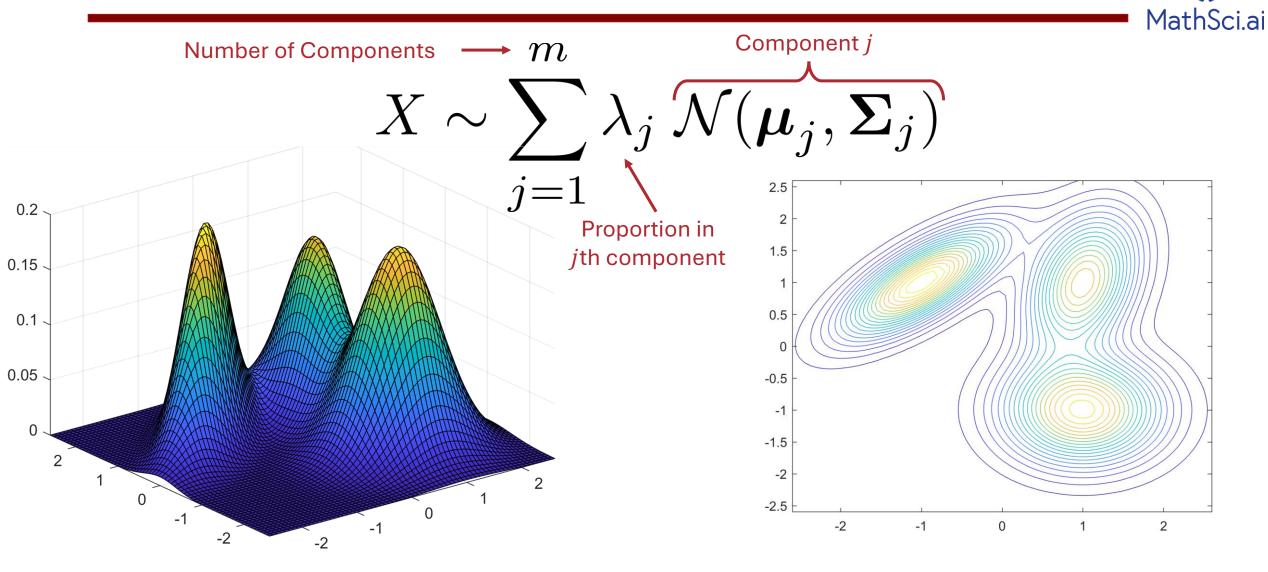
Kolda @ Texas A&M TRIPODS Distinguished Lecture

#### Gaussian Model of Data



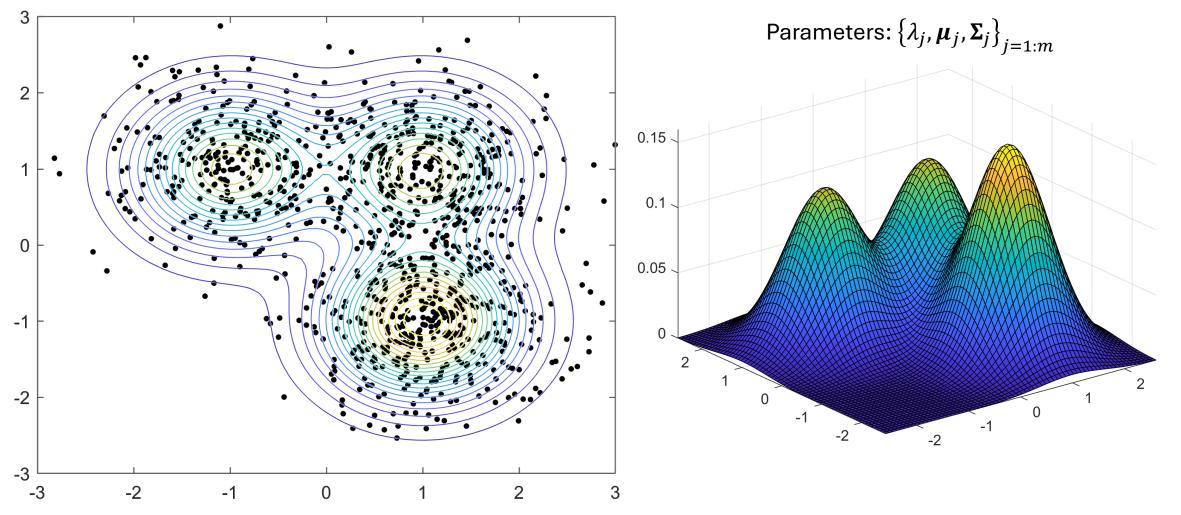


### Gaussian Mixture Model (GMM)



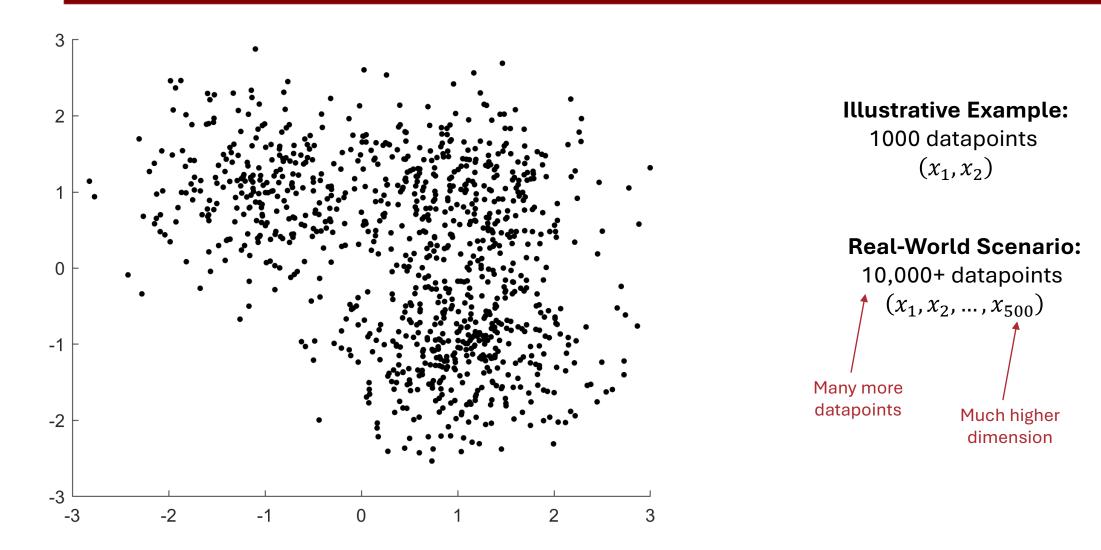
## GMM with Three Components





#### Motivation: Making Sense of Data





#### **Applications: Gaussian Mixtures**



#### 3 Density 2 Estimation 1 Clustering 0 -1 Anomaly -2 Detection -3

-3

-2

-1

0

1

3

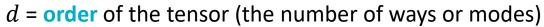
2

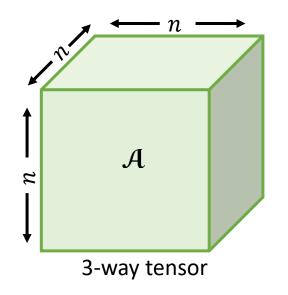


Tensors and Symmetric Tensors



### Tensors are Multi-dimensional Arrays





For this talk, all modes have the same size: n

 $n^d$  = number of entries for d-way tensor of dimension n

 $a_{iik} = (i, j, k)$  entry of 3-way tensor  $\mathcal{A}$ 

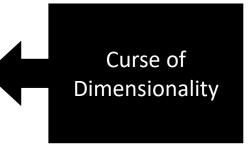
A tensor is **symmetric** if its entries are invariant under permutation, i.e.,

 $a_{ijk} = a_{ikj} = a_{jik} = a_{jik} = a_{jki} = a_{kij} = a_{kji}$ 

Curse of notation...

 $(i_1, i_2, ..., i_d) = index into tensor, i_k \in \{1, ..., n\}$  for k = 1, 2, ..., d

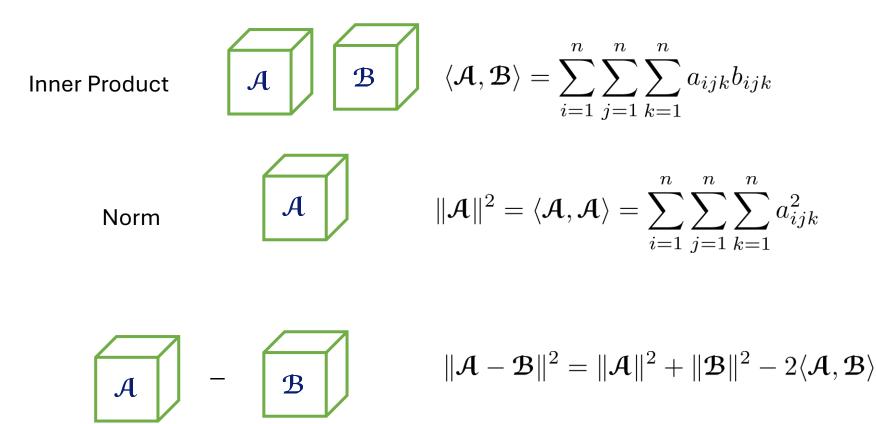




MathSci.ai

#### **Tensor Norm & Inner Product**

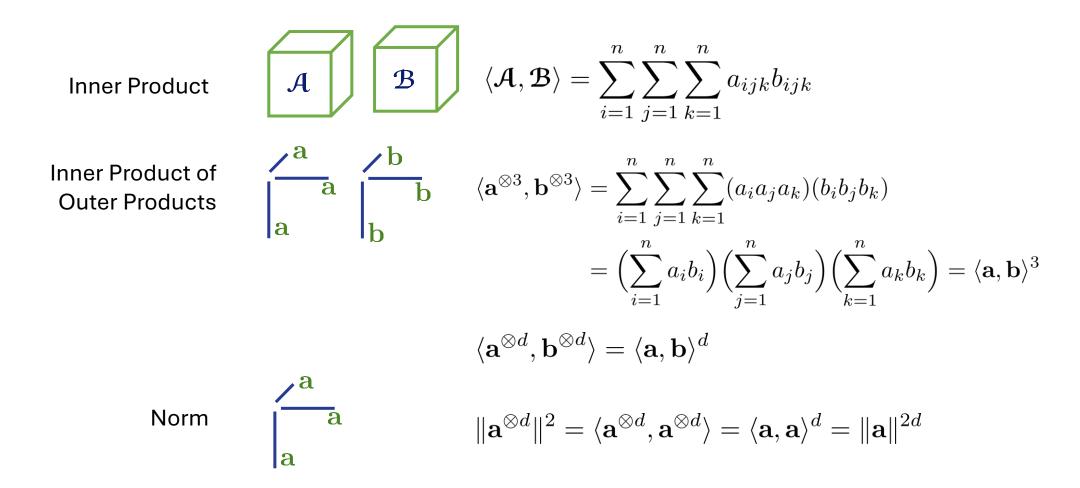




#### Symmetric Tensor Outer Product MathSci.ai $\mathbf{b} \in \mathbb{R}^n$ Visualization of 3-way $a_{ij} = b_i b_j$ $\mathbf{A} = \mathbf{b}^{\otimes 2} = \mathbf{b} \otimes \mathbf{b} \in \mathbb{R}^{n \times n}$ **Outer Product** $a_{ijk} = b_i b_j b_k$ = $\mathcal{A} = \mathbf{b}^{\otimes 3} = \mathbf{b} \otimes \mathbf{b} \otimes \mathbf{b} \in \mathbb{R}^{n \times n \times n}$ $\mathcal{A} = \mathbf{b}^{\otimes 4} = \mathbf{b} \otimes \mathbf{b} \otimes \mathbf{b} \otimes \mathbf{b} \in \mathbb{R}^{n \times n \times n \times n}$ $a_{ijk\ell} = b_i b_j b_k b_\ell$ $a_{i_1\cdots i_d}=b_{i_1}\cdots b_{i_d}$ $\mathcal{A} = \mathbf{b}^{\otimes d} = \mathbf{b} \otimes \cdots \otimes \mathbf{b} \in \mathbb{R}^{n \times \cdots \times n}$ d times

#### Outer Products, Inner Products, and Norms







# Approximate Method of Moments for Gaussian Mixtures

See Sherman & K (2020)

1/31/2022

Kolda @ Texas A&M TRIPODS Distinguished Lecture

#### Moments of a Multivariate Random Variable



Higher moments capture interactions between the variables

Random variable:  $X \in \mathbb{R}^n$ 

First moment:  $\mathbf{M}^{(1)} = \mathbb{E}(X)$   $\mathbf{M}^{(1)}(i) = \mathbb{E}(X_i)$ 

Second moment: 
$$\mathbf{M}^{(2)} = \mathbb{E}(X^{\otimes 2})$$
  $\mathbf{M}^{(2)}(i,j) = \mathbb{E}(X_i X_j)$ 

Third moment:  $\mathbf{M}^{(3)} = \mathbb{E}(X^{\otimes 3})$   $\mathbf{M}^{(3)}(i, j, k) = \mathbb{E}(X_i X_j X_k)$ 

dth moment:  $\mathbf{M}^{(d)} = \mathbb{E}(X^{\otimes d})$ 

*d*th moment is a symmetric tensor or order *d* 

#### Moments Define a Distribution



Higher moments capture interactions between the variables Random variable:  $X \in \mathbb{R}^n$ 

First moment:  $\mathbf{M}^{(1)} = \mathbb{E}(X)$ 

Second moment:  $\mathbf{M}^{(2)} = \mathbb{E}(X^{\otimes 2})$ 

Third moment:  $\mathbf{M}^{(3)} = \mathbb{E}(X^{\otimes 3})$ 

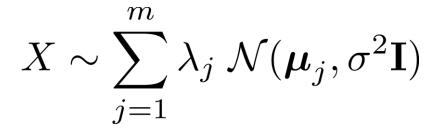
*d*th moment:  $\mathbf{M}^{(d)} = \mathbb{E}(X^{\otimes d})$ 

*"Method of Moments"* matches empirical and model moments to estimate the parameters of a distribution.

We focus primarily on matching just the dth moment

#### Gaussian Mixture Model: Small Spherical Covariance



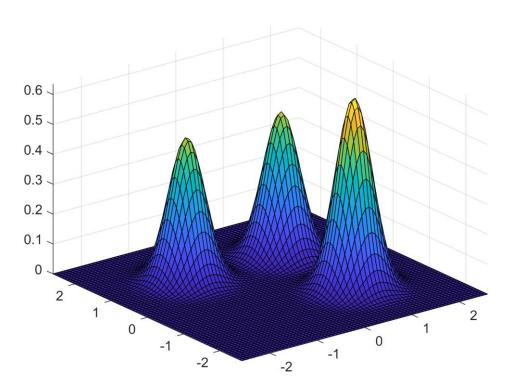


Moment of the Model:

$$\mathbf{\mathcal{M}}^{(d)} = \sum_{j=1}^{m} \lambda_j \boldsymbol{\mu}_j^{\otimes d} + \mathcal{O}(\sigma^2)$$

Approximate Moment of the Model:

$$\widetilde{\mathbf{\mathcal{M}}}^{(d)} = \sum_{j=1}^m \lambda_j \boldsymbol{\mu}_j^{\otimes d}$$



## Fitting the *d*th Moment

Given *r* observations:

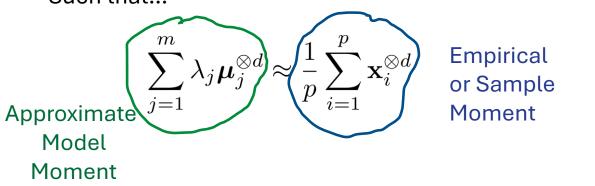
$$\{\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_p\} \subset \mathbb{R}^n$$

Find *m* weights and mean vectors:

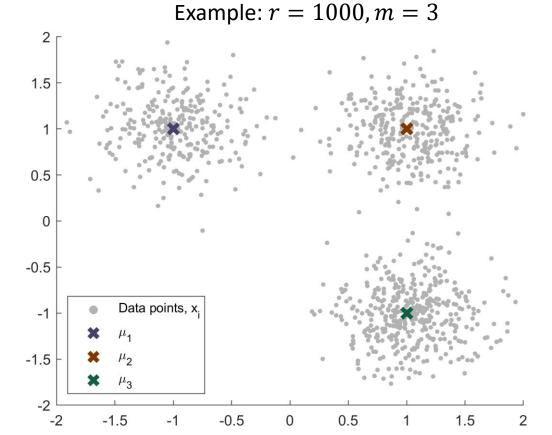
$$\{\lambda_1, \lambda_2, \dots, \lambda_m\} \subset \mathbb{R}$$
  
 $\{\mu_1, \mu_2, \dots, \mu_m\} \subset \mathbb{R}^n$ 

ignoring covariance σ<sup>2</sup>Ι

Such that...

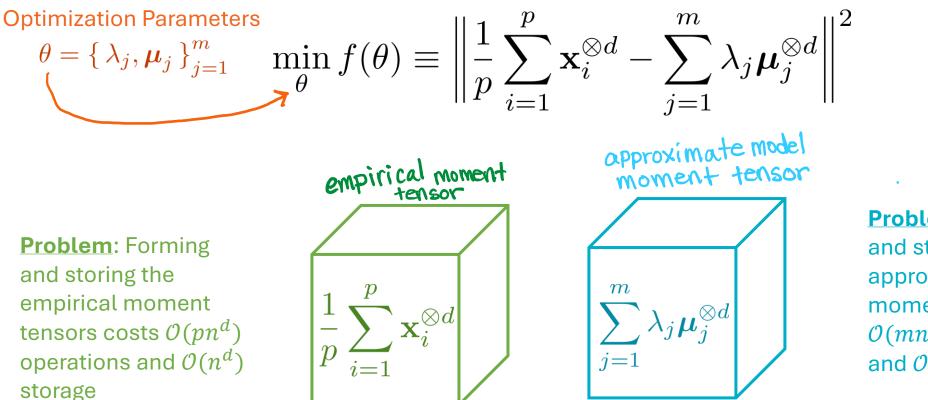


MathSci.ai



# MathSci.ai

# Optimization Formulation: Symmetric Tensor Decomposition



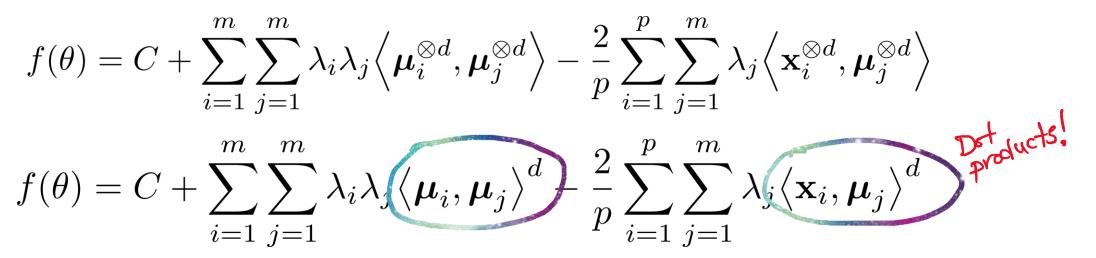
**Problem**: Forming and storing the approximate model moment tensor costs  $\mathcal{O}(mn^d)$  operations and  $\mathcal{O}(n^d)$  storage

<u>But</u> there is only  $\mathcal{O}(pn)$  data and  $\mathcal{O}(mn)$  parameters, so there is room for efficiency

#### **Optimization Formulation as Inner Products**

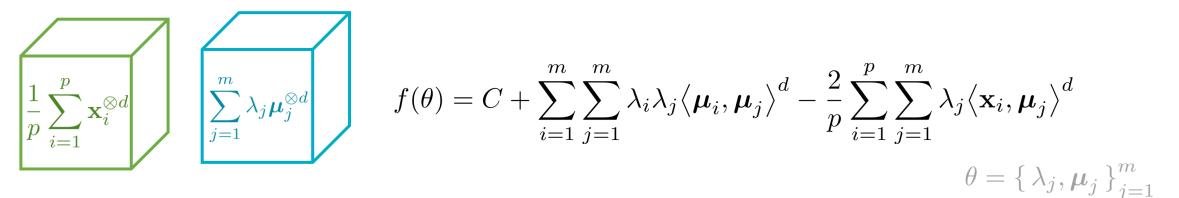


$$\min_{\theta} f(\theta) \equiv \left\| \frac{1}{p} \sum_{i=1}^{p} \mathbf{x}_{i}^{\otimes d} - \sum_{j=1}^{m} \lambda_{j} \boldsymbol{\mu}_{j}^{\otimes d} \right\|^{2} \qquad \theta = \{\lambda_{j}, \boldsymbol{\mu}_{j}\}_{j=1}^{m}$$
$$f(\theta) = \left\| \frac{1}{p} \sum_{i=1}^{p} \mathbf{x}_{i}^{\otimes d} \right\|^{2} + \left\| \sum_{j=1}^{m} \lambda_{j} \boldsymbol{\mu}_{j}^{\otimes d} \right\|^{2} - 2\left\langle \frac{1}{p} \sum_{i=1}^{p} \mathbf{x}_{i}^{\otimes d}, \sum_{j=1}^{m} \lambda_{j} \boldsymbol{\mu}_{j}^{\otimes d} \right\rangle$$
$$Constant$$



## **Casting as Optimization Problem**



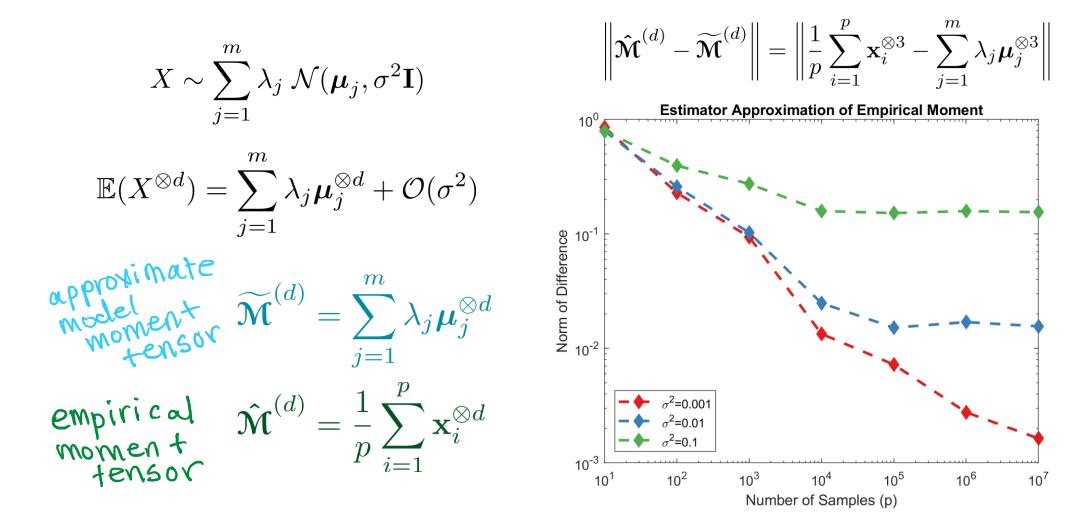


- Never need to form empirical or approximate model moments explicitly, overcoming curse of dimensionality
- Function can be calculated with only dot products, total work  $O(m^2n + pmn)$ and O(mn + pn) storage, versus  $O(mn^d + pn^d)$  as originally formulated
- Gradients equally efficient to calculate, via chain rule
- Issue: Inherent scaling problem (will come back to this later)
- Easy stochastic function and gradient if number of samples (p) is large...

$$\tilde{f}(\theta) = C + \sum_{i=1}^{m} \sum_{j=1}^{m} \lambda_i \lambda_j \langle \boldsymbol{\mu}_i, \boldsymbol{\mu}_j \rangle^d - \frac{2}{|\boldsymbol{\Omega}|} \sum_{i \in \boldsymbol{\Omega}} \sum_{j=1}^{m} \lambda_j \langle \mathbf{x}_i, \boldsymbol{\mu}_j \rangle^d$$

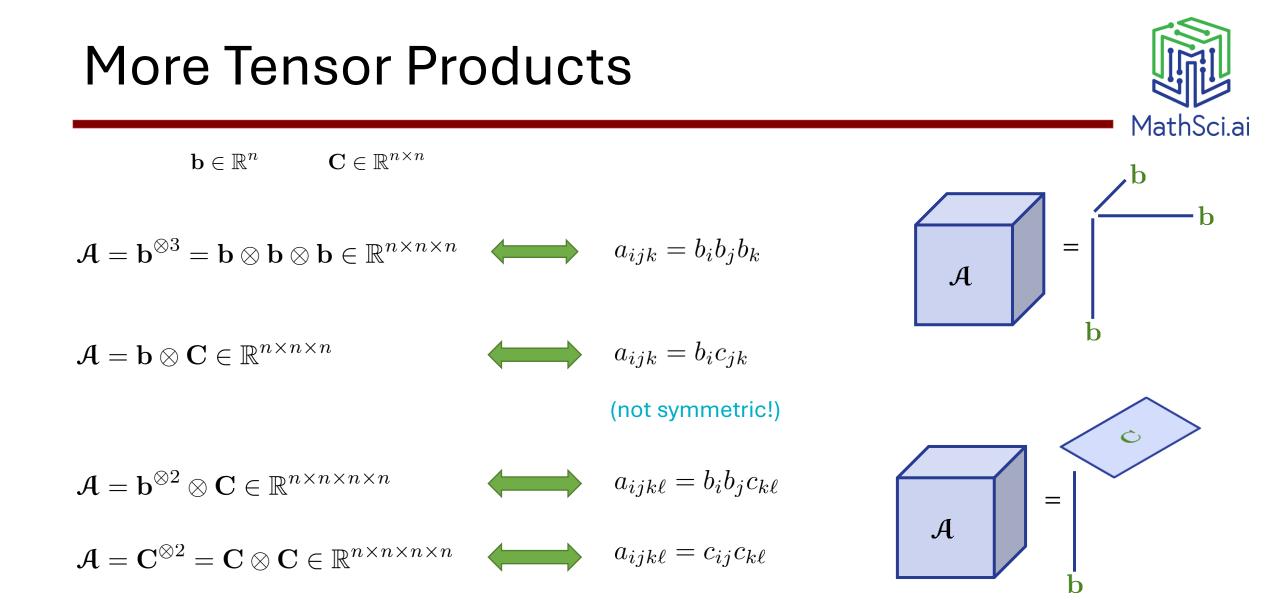
### But the Approximation is Biased







# More Tensors and Symmetric Tensors



#### Symmetrization



$$\operatorname{sym}\left(\operatorname{form}\right) = \frac{1}{6}\left(\operatorname{form} + \operatorname{form} + \operatorname{for$$

Lemma (Hackbusch, 2019)

$$\left\langle \operatorname{sym}(\mathcal{A}), \mathbf{b}^{\otimes d} \right\rangle = \left\langle \mathcal{A}, \mathbf{b}^{\otimes d} \right\rangle$$



#### Symmetric Tensors Correspond to Polynomials



 $\mathcal{A} \in \mathbb{R}^{n imes n imes n}$ 

$$\Phi[\mathcal{A}](z_1,\ldots,z_n) \equiv \sum_{i=1}^n \sum_{j=1}^n \sum_{k=1}^n a_{ijk} z_i z_j z_k = \langle \mathcal{A}, \mathbf{z}^{\otimes 3} \rangle$$

Proposition (Kileel-K-Pereira 2022)

$$\Phi[\mathbf{a}](z_1,\ldots,z_n) = \mathbf{a}^\mathsf{T}\mathbf{z}$$

 $\Phi[\mathbf{A}](z_1,\ldots,z_n) = \mathbf{z}^\mathsf{T} \mathbf{A} \mathbf{z}$ 

 $\Phi[\mathcal{A}](z_1,\ldots,z_n) = \Phi[\operatorname{sym}(\mathcal{A})](z_1,\ldots,z_n)$ 

 $\Phi[\mathcal{A} \otimes \mathcal{B}](z_1, \ldots, z_n) = \Phi[\mathcal{A}](z_1, \ldots, z_n) \cdot \Phi[\mathcal{B}](z_1, \ldots, z_n)$ 

**Binomial Theorem for Tensors** 

(Kileel-K-Pereira 2022)

$$(\mathbf{a} + \mathbf{b})^{\otimes d} = \sum_{k=0}^{d} {d \choose k} \operatorname{sym} \left( \mathbf{a}^{\otimes k} \otimes \mathbf{b}^{\otimes d-k} \right)$$

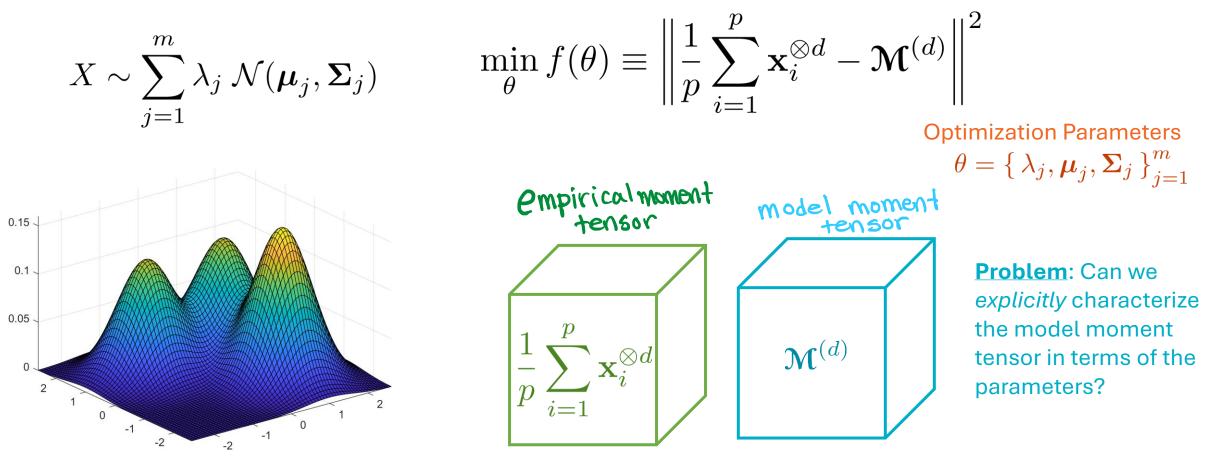


## Method of Moments for Gaussian Mixture Models – General Scenario

1/31/2022

#### Gaussian Mixture Model: General Case

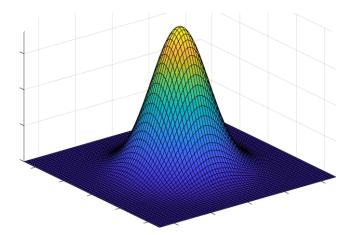




#### Gaussian Model Moment



Let 
$$X \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}), \, \boldsymbol{\mathcal{M}}^{(d)} \equiv \mathbb{E}(X^{\otimes d})$$
. Then



$$\begin{split} \mathfrak{M}^{(1)} &= \boldsymbol{\mu} \\ \mathfrak{M}^{(2)} &= \boldsymbol{\mu}^{\otimes 2} + \boldsymbol{\Sigma} \\ \mathfrak{M}^{(3)} &= \boldsymbol{\mu}^{\otimes 3} + 3 \operatorname{sym}(\boldsymbol{\mu} \otimes \boldsymbol{\Sigma}) \\ \mathfrak{M}^{(4)} &= \boldsymbol{\mu}^{\otimes 4} + 6 \operatorname{sym}(\boldsymbol{\mu}^{\otimes 2} \otimes \boldsymbol{\Sigma}) + 3 \operatorname{sym}(\boldsymbol{\Sigma}^{\otimes 2}) \end{split}$$

#### **Theorem** (Kileel-K-Pereira 2022)

$$\mathbf{\mathcal{M}}^{(d)} = \sum_{k=0}^{\lfloor d/2 \rfloor} {\binom{d}{2k}} \frac{2k!}{k! 2^k} \operatorname{sym} \left( \boldsymbol{\mu}^{\otimes d-2k} \otimes \boldsymbol{\Sigma}^{\otimes k} \right)$$

#### Proof techniques

- Equivalence of symmetric tensors and polynomials
- Marginals of multivariate Gaussians are univariate Gaussian
- Binomial theorem

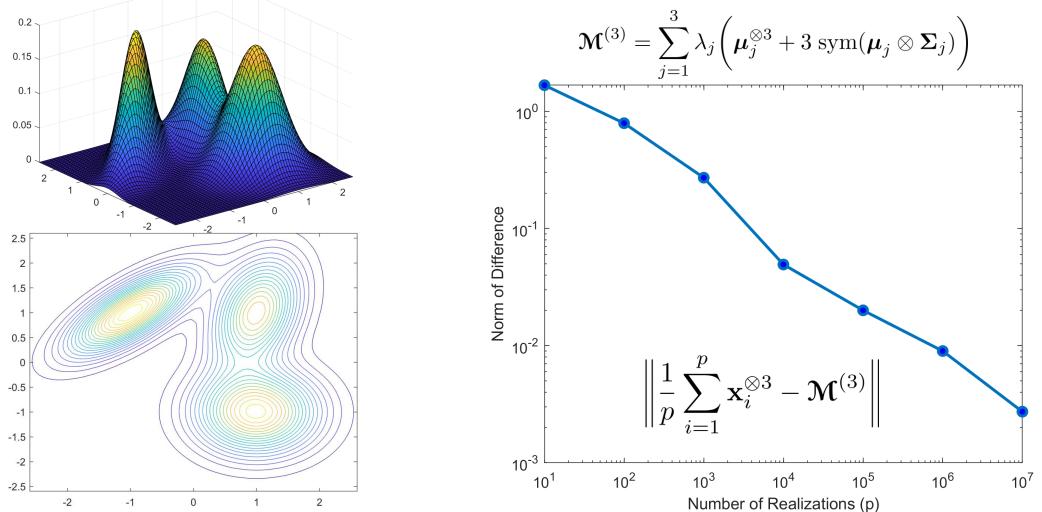


#### **Theorem** (Kileel-K-Pereira 2022)

Let 
$$X \sim \sum_{j=1}^{m} \lambda_j \mathcal{N}(\boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j), \, \mathfrak{M}^{(d)} \equiv \mathbb{E}(X^{\otimes d}).$$
  
Then  
 $\mathfrak{M}^{(d)} = \sum_{j=1}^{m} \lambda_j \mathfrak{M}_j^{(d)}$   
 $\mathfrak{M}_j^{(d)} = \sum_{k=0}^{m} \lambda_j \mathfrak{M}_j^{(d)}$   
 $\mathfrak{M}_j^{(d)} = \sum_{k=0}^{\lfloor d/2 \rfloor} \binom{d}{2k} \frac{2k!}{k! 2^k} \operatorname{sym}\left(\boldsymbol{\mu}_j^{\otimes d-2k} \otimes \boldsymbol{\Sigma}_j^{\otimes k}\right)$ 

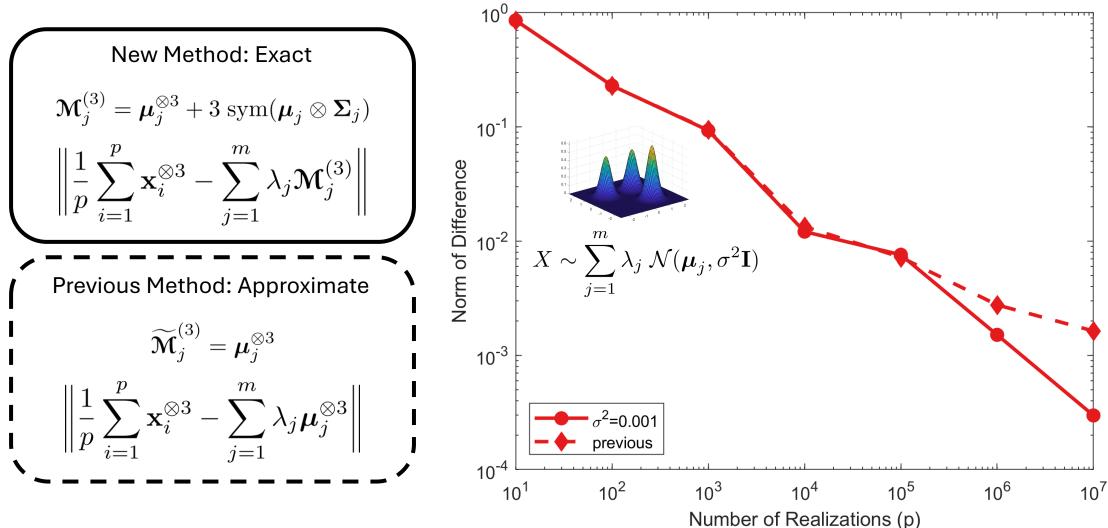
#### **Empirical versus GMM Model Moment**





### Previous Approach was Biased

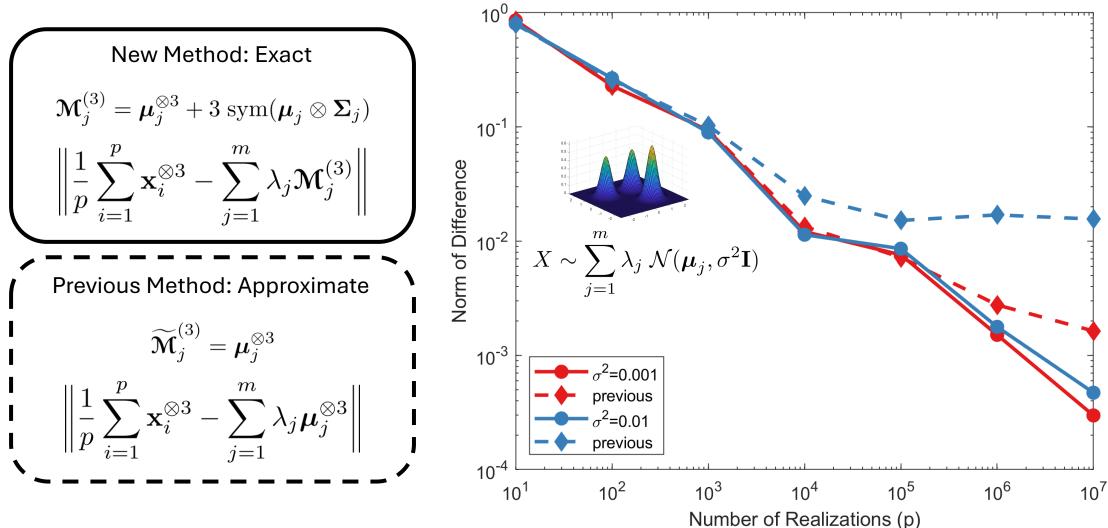




Kolda @ Texas A&M TRIPODS Distinguished Lecture

# Previous Approach was Biased

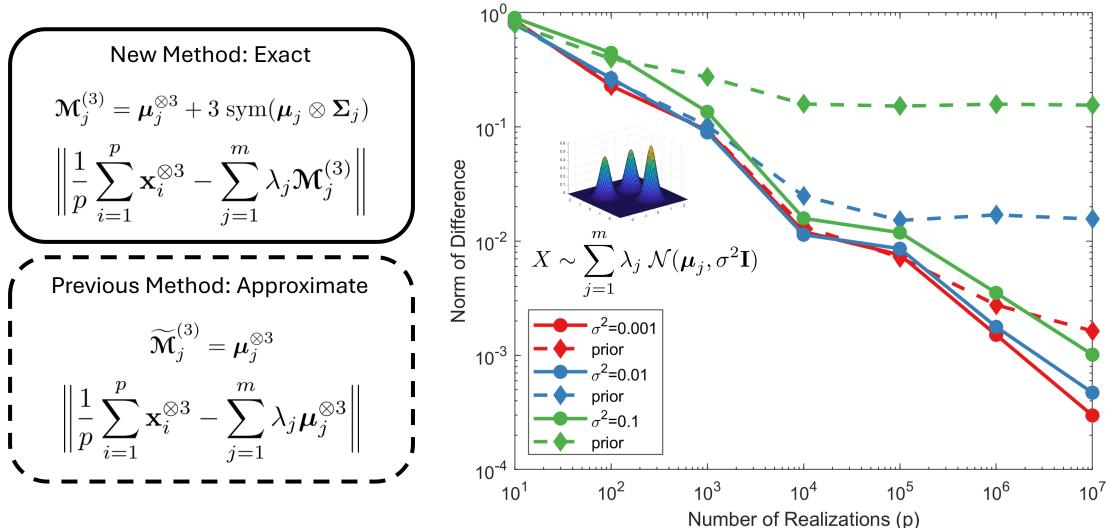




Kolda @ Texas A&M TRIPODS Distinguished Lecture

# Previous Approach was Biased

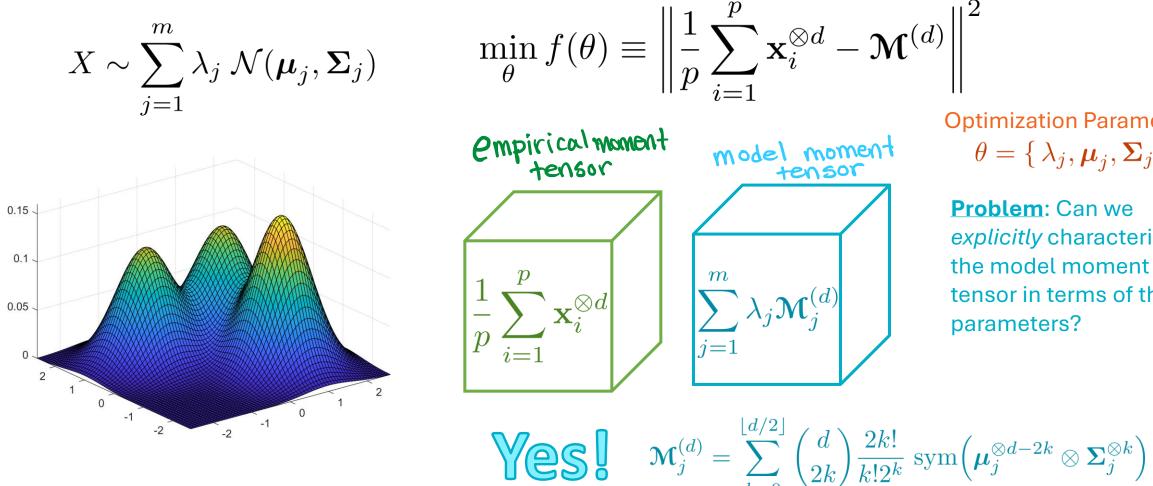




Kolda @ Texas A&M TRIPODS Distinguished Lecture

# **Optimization Formulation**





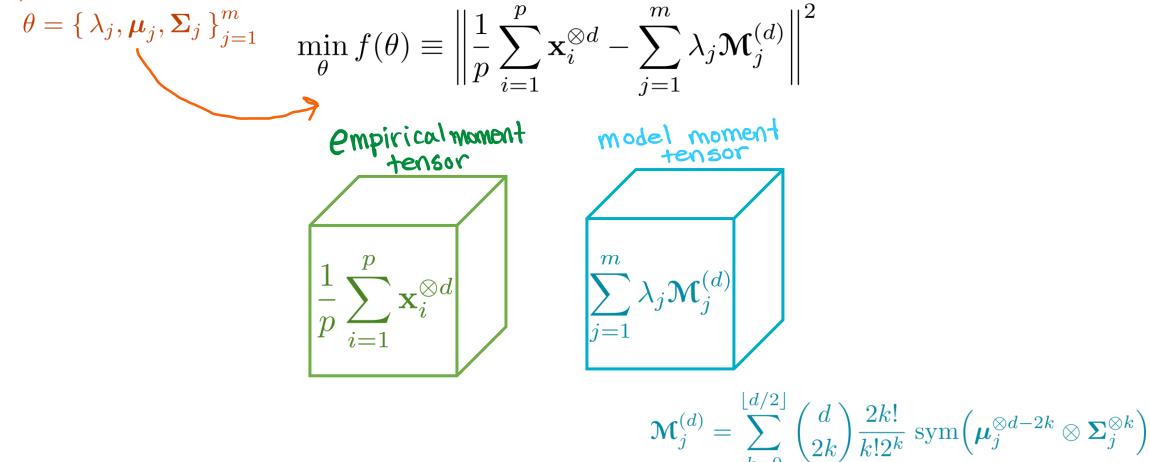
**Optimization Parameters**  $\theta = \left\{ \lambda_j, \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j \right\}_{j=1}^m$ 

Problem: Can we explicitly characterize the model moment tensor in terms of the parameters?

# **Optimization Formulation**



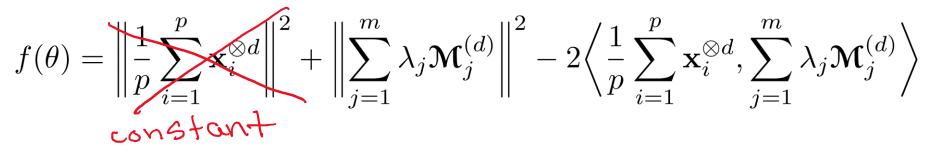
**Optimization Parameters** 



### **Optimization Formulation as Inner Products**



$$\min_{\theta} f(\theta) \equiv \left\| \frac{1}{p} \sum_{i=1}^{p} \mathbf{x}_{i}^{\otimes d} - \sum_{j=1}^{m} \lambda_{j} \mathbf{\mathcal{M}}_{j}^{(d)} \right\|^{2}$$



### Example Calculation: d = 3



$$f(\theta) = C + \sum_{i=1}^{m} \sum_{j=1}^{m} \lambda_i \lambda_j \left\langle \mathbf{M}_i^{(d)}, \mathbf{M}_j^{(d)} \right\rangle - \frac{2}{p} \sum_{i=1}^{p} \sum_{j=1}^{m} \lambda_j \left\langle \mathbf{x}_i^{\otimes d}, \mathbf{M}_j^{(d)} \right\rangle$$

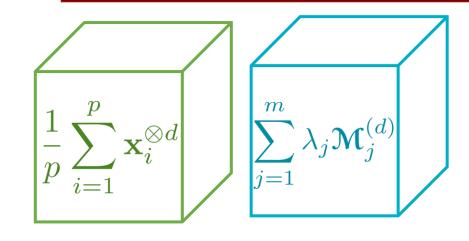
$$\mathbf{\mathcal{M}}_{j}^{(3)} = \mathbf{\mu}_{j}^{\otimes 3} + 3 \operatorname{sym}(\mathbf{\mu}_{j} \otimes \mathbf{\Sigma}_{j})$$

$$\begin{aligned} \langle \mathbf{x}_i^{\otimes 3}, \mathbf{\mathcal{M}}_j^{(3)} \rangle &= \left\langle \mathbf{x}_i^{\otimes 3}, \boldsymbol{\mu}_j^{\otimes 3} \right\rangle + 3 \left\langle \mathbf{x}_i^{\otimes 3}, \operatorname{sym}(\boldsymbol{\mu}_j \otimes \boldsymbol{\Sigma}_j) \right\rangle \\ &= \left( \mathbf{x}_i^{\mathsf{T}} \boldsymbol{\mu}_j \right)^3 + 3 \left\langle \mathbf{x}_i^{\otimes 3}, \boldsymbol{\mu}_j \otimes \boldsymbol{\Sigma}_j \right\rangle \\ &= \left( \mathbf{x}_i^{\mathsf{T}} \boldsymbol{\mu}_j \right)^3 + 3 \left( \mathbf{x}_i^{\mathsf{T}} \boldsymbol{\mu}_j \right) \left( \mathbf{x}_i^{\mathsf{T}} \boldsymbol{\Sigma}_j \mathbf{x}_i \right) \end{aligned}$$

Theory for computing inner product of two GMM moment tenors or empirical and GMM moment tensor – see forthcoming arXiv paper for details!

# **Casting as Optimization Problem**





$$\min_{\theta} f(\theta) = C + \sum_{i=1}^{m} \sum_{j=1}^{m} \lambda_i \lambda_j \left\langle \mathbf{M}_i^{(d)}, \mathbf{M}_j^{(d)} \right\rangle - \frac{2}{p} \sum_{i=1}^{p} \sum_{j=1}^{m} \lambda_j \left\langle \mathbf{x}_i^{\otimes d}, \mathbf{M}_j^{(d)} \right\rangle$$
$$\theta = \left\{ \lambda_j, \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j \right\}_{j=1}^{m}$$

- Never need to form empirical or model moments explicitly, overcoming curse of dimensionality
- Function can be calculated using simple calculations, total work  $O(m^2n + pmn^2 + m^2n^3)$  per iteration and O(mn + pn) storage
- Gradients can be calculated as well
- Easy stochastic function and gradient if number of samples (p) is large, as before
- Issue: Inherent scaling problem

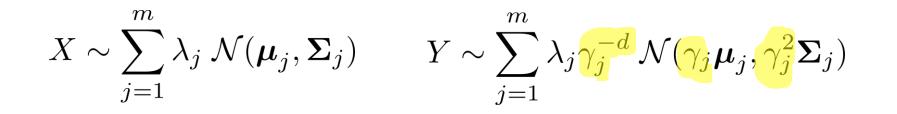


$$X \sim 0.5 \,\mathcal{N}(\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1) + 0.5 \,\mathcal{N}(\boldsymbol{\mu}_2, \boldsymbol{\Sigma}_2)$$
$$Y \sim 0.25 \,\mathcal{N}\left((\frac{1}{2})^{1/3} \boldsymbol{\mu}_1, (\frac{1}{2})^{2/3} \boldsymbol{\Sigma}_1\right) + 0.75 \,\mathcal{N}\left((\frac{3}{2})^{1/3} \boldsymbol{\mu}_2, (\frac{3}{2})^{2/3} \boldsymbol{\Sigma}_2\right)$$

$$\mathbb{E}(X^{\otimes 3}) = \mathbb{E}(Y^{\otimes 3})$$

## Non-uniqueness problem





**Problem**: Moments of order *d* are the same!

**FIX**: Append a constant *c* to the end of every observation vector, creating vectors of dimension n + 1

**RESULT**: *Implicitly,* a weighted combination of *all the moments* from 1 to d. This means we include all moments up to order d in the optimization.

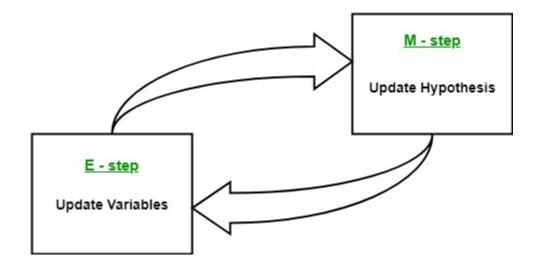
See forthcoming arXiv posting for full details!!



# Numerical Results and Conclusions

### State of the Art: Expectation Maximization (EM)





<u>EM Algorithm in a Nutshell</u> Make initial guesses for parameters Repeat until log-likelihood converges:

- 1. Compute membership weights for each datapoint
- 2. Update the component parameters using the membership weights

#### EM is State of the Art

- Inexpensive
- Relatively easy to implement
- Optimizing a different cost function
- Sensitive to initialization
- Sensitive to overlapping Gaussians

MoM has theoretical advantages but has not been used much in practice previously because of its great expense

See, e.g., Xu and Jordan (1996) for discussion of its robustness

# Method of Moments can beat EM



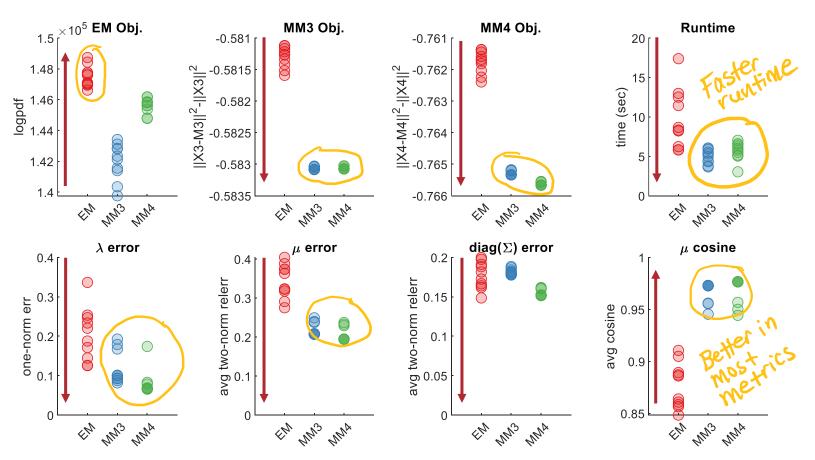
- 7.05 ſ<sup>×10<sup>5</sup> EM Obj.</sup> MM3 Obj. MM4 Obj. Runtime -0.237 -0.188  $\Theta$ time (sec)  $\mathbb{C}$  $\bigcirc$ Jpdbol 6.95 6.9  $\bigcirc$ 6.85 2 6.8 -0.2395 -0.192 MMS NMS MMA NNNS NMS NMA MMA MNA EM EM EM C/N  $diag(\Sigma)$  error  $\lambda$  error  $\mu$  error u cosine 0.3<sub>T</sub> 0.25 0.4 avg two-norm relerr .0 0 0 0.2 0.15 0.1 0.1 0.1 0.2 oue-norm err 0.3 avg cosine 0.9  $\bigcirc$ 0.2 0.85 EM MMS MMS NNNS MMA MMS MMA EM MMA MNA EM EN
- Randomly-generated problem with overlapping Gaussians
  - Diagonal covariances ٠
  - Dimensionality: n = 100٠
  - Number of Gaussians: m = 20٠
  - Observations: p = 8000٠
- Compared three methods •
  - **EM: Expectation Maximization** •
  - MM3: Method of Moments, d = 3٠
  - MM4: Method of Moments, d = 4٠
- 10 runs each with different initial guesses

# Method of Moments can beat EM



- Randomly-generated problem with overlapping Gaussians
  - Diagonal covariances
  - Dimensionality: n = 100
  - Number of Gaussians: m = 20
  - Observations: p = 8000
- Compared three methods
  - EM: Expectation Maximization
  - MM3: Method of Moments, d = 3
  - MM4: Method of Moments, d = 4
- 10 runs each with different initial guesses

#### Same setup as previous slide except higher noise

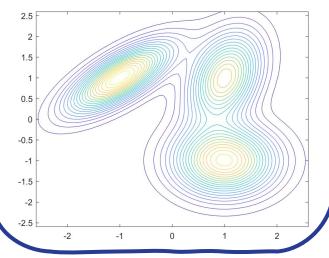


# **Related Works**



Key Differences in Our Work

- Novel tensor formulation of Gaussian moments
- No spherical or axisaligned covariance assumptions
- Computationally efficient, no exponential dependence on *d*



- Many theoretical advantages to method of moments and connections to tensors
  - Hsu and Kakade (2013) diagonal covariance,  $d \leq 3$
  - Ge, Huang, Kakade (2015) vectorized covariances, loses symmetries
  - Bakshi, Diakonikolas, Jia, Kane, Kothari, and Vempala (2020) robust learning using tensor decomposition
  - Khouga, Mattei, Mourrian (2021) GMM identifiability
- Computational approaches (limited handling of covariances)
  - Anandkumar, Ge, Hsu, Kakade (2014) and Anandkumar, Ge, Hsu, Kakade, Telegarsky (2014) – orthogonal symmetric tensor decomposition
  - Sherman & K., 2020 (general) symmetric tensor decomposition, emphasis of implicit computation to avoid curse of dimensionality
- Inner products of moment tensors
  - Muandet, Fukumizu, Dinuzzo, Schölkopf (2012) up to 3rd order

# Take-aways and Future Work

- Our focus: Method of moments for Gaussian mixture models (GMMs)
- Key results
  - Formulation of GMM moment in terms of tensor outer products

$$\mathbf{\mathcal{M}}^{(d)} = \sum_{j=1}^{m} \sum_{k=0}^{\lfloor d/2 \rfloor} {d \choose 2k} \frac{2k!}{k! 2^k} \lambda_j \operatorname{sym} \left( \boldsymbol{\mu}_j^{\otimes d-2k} \otimes \boldsymbol{\Sigma}_j^{\otimes k} \right)$$

- Efficient computation and storage, avoiding exponential dependence on moment order
- Novel approach to scaling ambiguity using augmentation
- Amenable to stochastic formulations
- Plus...dot product of moment tensors in terms of Bell polynomials, avoiding exponential dependence on moment order
- Plus...modifying empirical moment tensor to "remove" Gaussian noise
- Future work
  - Implementation details, especially for general Gaussians
  - Analysis of optimization landscape and comparison to that of max likelihood
  - Bounding number of samples required for accurate estimation
  - Application studies

Tensor Moments of Gaussian Mixture

Models: Theory and Applications

J. Kileel, T. Kolda, and J. M. Pereira.



MathSci.ai



```
JOE KILEEL, TAMARA G. KOLDA, AND JOÃO M. PEREIRA
```

ANSTACT. Gaussian mixture models (GMM) are fundamental tools in statistical and data sciences. We study the moments of multivariate Gaussians and GMMs. The dth moment of an n-dimensional random variable is a symmetrie d-way tensor of size  $n^4$ , so working with moments is assumed to be prohibitively expensive for d > 2 and larger values of n. In this work, we develop theory and numerical methods for implicit compatiations with moment adjust expensions for the momenta of the structure of the symmetry of the space of the symmetry of the symmetry of the symmetry of the adjust expensions of the momenta in terms of symmetrized tensor products, relying on the correspondence between symmetric tensors and homogeneous et of observations, which can be formulated as a moment-matching optimization problem. If there is a known and commate covariance matrix, then it is work also has relevance to symmetric tensor decomposition. This work also has relevance to implicit computations on the momento Gaussians. Momerical results liburate the mumerical difficurely of these approaches.

#### 1. INTRODUCTION

The Gaussian mixture model (GMM) is a fundamental tool in statistical and data sciences. The Gaussian distribution, also known as the normal distribution, is the limiting distribution of any sum of random variables. A finite convex combination (i.e., a mixture) of Gaussians is a GMM. Utilization of GMMs is ubiquitous in density approximation, clustering, and anomaly detection, finding application in domains such as image processing, biomedicine, financial forecasting, text analytics, process monitoring, and much mere.

In this work, we consider the characterization of the moments of multivariate GMMs, with the primary aim of determining the parameters of GMM by matching sample and model moments. In contrast to the well-known expectation maximization (EM) method that does maximizmum likelihood estimation (Hastie et al., 2000; Murphy, 2012; Xu and Jordan, 1990), moment matching may have superior theoretical properties (Lindasy and Basak, 1993; Høu and Kakade, 2013; Ge et al., 2015; Baskhi et al., 2021; Khouja et al., 2021; Kane, 2021).

The main difficulty with moments is that a *d*th-order moment can be prohibitive to compute and store since it involves the expectations of powers of a random variable; in other words, the *d*th moment of an *n*-dimensional random variable is a

Date: January 24, 2022. 2020 Mathematics Subject Classification. Primary

0.15

0.1

0.05